

Problem 1. If c is a constant such that $X_n \xrightarrow{\mathcal{L}} c$, prove that $X_n \xrightarrow{P} c$.

Problem 2. Suppose X_1, X_2, \dots are iid Bernoulli(p) random variables. Let $S_n = \sum_{i=2}^n X_i X_{i-1}$. Find (with justification) the asymptotic distribution of S_n/n .

Problem 3. Either construct an example of each of the following or prove that no such example exists:

(a) [1 point] A sequence X_1, X_2, \dots of random variables, a random variable X , and a constant x_0 such that $X_n \xrightarrow{\mathcal{L}} X$ and $F_n(x_0) \not\rightarrow F(x_0)$.

(b) [1 point] A sequence X_1, X_2, \dots of random variables such that $X_n \xrightarrow{P} 0$ and $E(X_n) \rightarrow 1$.

Problem 4. Let X_1, \dots, X_n be iid from a continuous symmetric distribution centered at 0. Suppose (Y_1, \dots, Y_n) is a permutation of (X_1, \dots, X_n) satisfying $|Y_1| < |Y_2| < \dots < |Y_n|$; that is, the Y_i are the X_i arranged in order of increasing absolute value.

Let $W_n = \sum_{i=1}^n iI\{Y_i > 0\}$ be the usual signed-rank statistic. Derive the asymptotic distribution of W_n , justifying your steps.

[We have seen at least two ways to do this. You may of course use any valid method you choose.]

Problem 5. Suppose X_1, \dots, X_n are iid random variables with cdf $F(x)$. Let \tilde{X}_n denote the sample median. Suppose we wish to estimate $h(F) = \text{Var}(\tilde{X}_n) < \infty$. We use a bootstrap scheme in which we draw B random samples of size n from \hat{F}_n , the empirical cdf, and let M_i be the sample median of the i th sample, $i = 1, \dots, B$.

If \bar{M}_B denotes $(1/B) \sum_{i=1}^B M_i$, explain (with justification) what happens to

$$\frac{1}{B} \sum_{i=1}^B (M_i - \bar{M}_B)^2$$

as $B \rightarrow \infty$.

Problem 6. Suppose X_1, X_2, \dots are iid Poisson(θ) random variables. Find the asymptotic distribution of $(S_n - E S_n)/\sqrt{\text{Var } S_n}$, where

$$S_n = \sum_{i=2}^{n+1} X_i I\{X_{i-1} = 0\}.$$

The Poisson(θ) distribution has expectation θ , variance θ , and mass function $p(x) = e^{-\theta}\theta^x/x!$ for x a nonnegative integer.

Problem 7. Suppose that X_1, \dots, X_n are iid with

$$P(X_i = 0) = \theta \quad \text{and} \quad P(X_i = -\sqrt{1-\theta}) = P(X_i = \sqrt{1-\theta}) = \frac{1-\theta}{2}.$$

Define $Y_i = I\{X_i = 0\}$.

Let $\underline{Z}^{(i)} = (X_i, Y_i)$. Find a function $g(x, y) = [g_1(x, y), g_2(x, y)]$ that is a variance-stabilizing transformation in the sense that the asymptotic distribution of $g[\bar{Z}] - g[E(\underline{Z}^{(i)})]$ has a covariance matrix that doesn't depend on θ , where \bar{Z} is the sample mean of the $\underline{Z}^{(i)}$.

You may find it helpful to know that

$$\frac{d}{dt} 2 \sin^{-1}(\sqrt{t}) = \frac{1}{\sqrt{t(1-t)}}.$$

Problem 8. Let X_1, \dots, X_n be an iid sample from $\text{Beta}(\alpha, 1)$; that is, $f_\alpha(x) = \alpha x^{\alpha-1}$. You may assume without proof that all relevant regularity conditions apply to the beta distribution.

The $\text{Beta}(\alpha, \beta)$ distribution has expectation $\alpha/(\alpha + \beta)$, variance $\alpha\beta/[(1 + \alpha + \beta)(\alpha + \beta)^2]$, and density $\Gamma(\alpha + \beta)x^{\alpha-1}(1-x)^{\beta-1}I\{0 < x < 1\}/[\Gamma(\alpha)\Gamma(\beta)]$.

(a) Compute a Wald test statistic W_n that is asymptotically standard normal under $H_0 : \alpha = \alpha_0$.

(b) Compute a Rao score test statistic R_n that is asymptotically standard normal under $H_0 : \alpha = \alpha_0$.

Problem 9. Suppose X_1, X_2, \dots are independent with $X_i \sim \text{Beta}(\alpha_i, \alpha_i)$, where $0 < \alpha_i < 2$. Prove that

$$\frac{\sum_{i=1}^n (X_i - \frac{1}{2})}{\sqrt{s_n^2}} \xrightarrow{\mathcal{L}} N(0, 1),$$

where $s_n^2 = \sum_{i=1}^n \text{Var}(X_i)$, by verifying the Lindeberg condition or the Lyapunov condition.

The $\text{Beta}(\alpha, \beta)$ distribution has expectation $\alpha/(\alpha + \beta)$, variance $\alpha\beta/[(1 + \alpha + \beta)(\alpha + \beta)^2]$, and density $\Gamma(\alpha + \beta)x^{\alpha-1}(1-x)^{\beta-1}I\{0 < x < 1\}/[\Gamma(\alpha)\Gamma(\beta)]$.

Problem 10. Let X_1, \dots, X_n be an iid sample from $\text{Poisson}(\theta)$. Throughout this problem, you may assume that the Poisson distribution satisfies all relevant regularity conditions.

(a) Show that the Jeffreys prior on $(0, \infty)$ is the improper prior density $\lambda(\theta) = 1/\sqrt{\theta}$. To do this, it suffices to show that $\lambda(\theta)$ is proportional to $\sqrt{I(\theta)}$.

The $\text{Poisson}(\theta)$ distribution has expectation θ , variance θ , and mass function $p(x) = e^{-\theta}\theta^x/x!$ for x a nonnegative integer.

(b) Show that with the improper Jeffreys prior $\lambda(\theta) = 1/\sqrt{\theta}$, the posterior distribution of θ is gamma. Find the Bayes estimator $\delta_n = E(\theta | X_1, \dots, X_n)$ using this prior. Finally, give the asymptotic distribution of $\sqrt{n}(\delta_n - \theta)$.

The $\text{Gamma}(\alpha, \beta)$ distribution has expectation α/β , variance α/β^2 , and density $\beta^\alpha x^{\alpha-1}e^{-\beta x}I\{x > 0\}/\Gamma(\alpha)$.